

Strategies for Using Propensity Scores Well

A Workshop given by

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**Center for Health Care Research and Policy
Case Western Reserve University**

at the

**6th International Conference for
Health Policy Research**

**Friday, October 28, 2005
10:30 AM to 12:15 PM**

Boston Park Plaza Hotel and Towers

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Center for Health Care Research & Policy

Case Western Reserve University / MetroHealth Medical Center



Strategies for Using Propensity Scores Well

6th International Conference for Health Policy
Research – Boston – October 28, 2005

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Propensity Scores in the News!

Firearm Violence Exposure and Serious Violent Behavior

Jeffrey B. Bingenheimer,^{1*} Robert T. Brennan,² Felton J. Earls²

To estimate the cause-effect relationship between exposure to firearm violence and subsequent perpetration of serious violence, we applied the analytic method of propensity stratification to longitudinal data on adolescents residing in Chicago, Illinois. Results indicate that exposure to firearm violence approximately doubles the probability that an adolescent will perpetrate serious violence over the subsequent 2 years.

Bingenheimer et al. (2005), Holden (2005)

Acknowledgments

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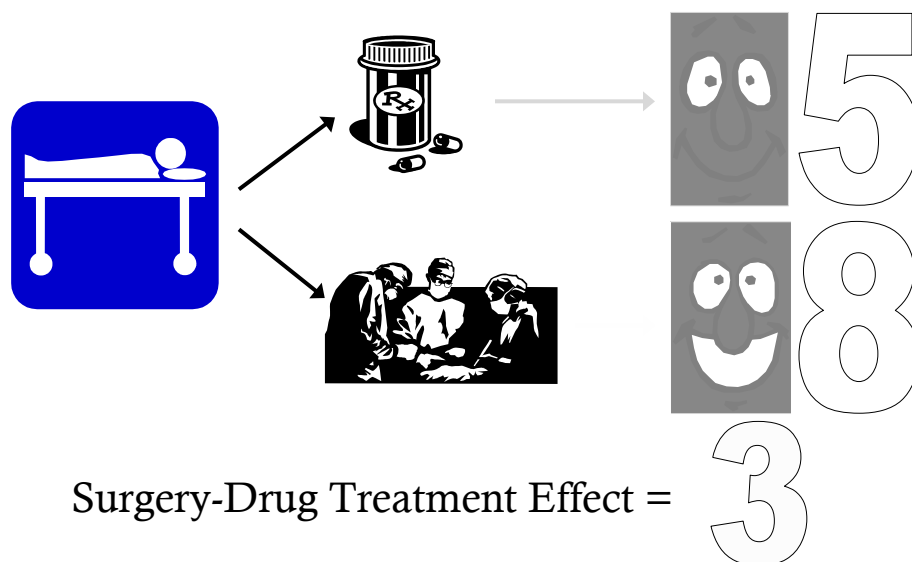
Workshop Outline

1. A Brief Review of Propensity Methods
2. Estimating the Propensity Score Effectively
3. Assessing and Displaying Covariate Balance
4. Choosing Analytic Techniques
5. Communicating Results to a Non-Statistical Audience

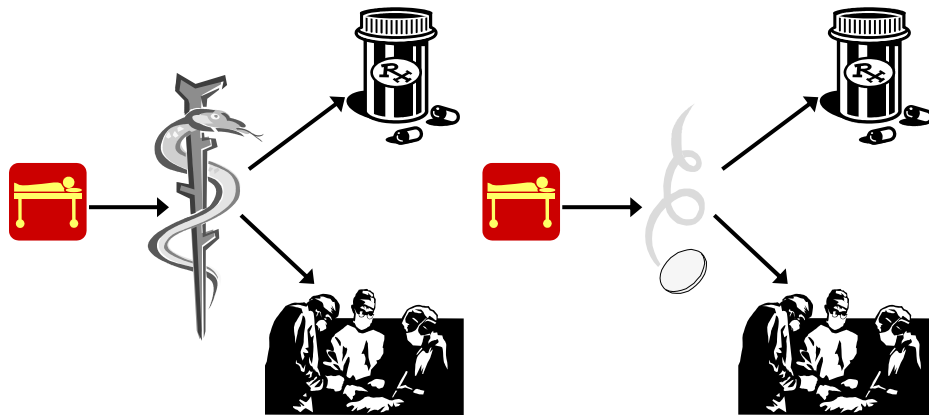
Part One

A (Brief) Review of Propensity Score Methods

Causal Treatment (Exposure) Effects
in terms of Potential Outcomes



Observational vs. Randomized Studies



In observational studies,
the researcher
does not randomly
allocate the treatments.

Randomization ensures
that subjects receiving
different treatments
are comparable.

Some Advantages of Smart Observational Studies

- Address chief criticism of RCTs - limited generalizability / external validity
- Data are widely (increasingly) available
 - Often reduced cost and time to get answer
- Enable examination of exposure in “real life”
- May enable examination of effect size and “entrenched practices”
- Large sizes permit investigation of exposures with smaller effect sizes

The Database You Wish You Had Vs. Harsh Reality

Patient	Outcome under Treatment	Outcome under No Treatment	Treatment's Impact
A	12	?	?
B	7	?	?
C	?	3	?
D	?	9	?

Assessing Causal Effect of Exposure on Outcome

- Objective: Draw causal inferences between [use of treatment vs. control] and outcome
- Standard Approach: Risk Adjustment
 - Problem: Selection Bias (people getting treatment are different from people getting control in ways that affect outcome)
- Idea: Compare treated to control subjects that looked similar (had similar propensity for treatment) prior to the treatment decision

Rubin Causal Model Framework

- Potential Outcomes posits two random variables to describe the outcome for each subject in the study
 - Y_T = Outcome under the treatment condition
 - Y_C = Outcome under the control condition
 - $\mathbf{X}$ = vector of covariates
 - Z = treatment assignment variable (1 if treated, 0 if control)

Rubin (1997), Rosenbaum (2002)

Making Inference About Treatment Effects Possible

- Strongly Ignorable Treatment Assignment Assumption:
 - The outcome variables Y_T and Y_C are assumed to be conditionally independent of the treatment assignment variable Z , given the covariates $\mathbf{X}$
 - This is a “no hidden bias” assumption
 - We assume that $\mathbf{X}$ contains all relevant information about treatment assignment

The Propensity Score

Pr(treatment given covariates)

- Definition: The conditional probability of receiving a given exposure (treatment) given a vector of measured covariates.
- Reduces baseline information to a single composite summary of the covariates.

$$\ln\left(\frac{PS}{1-PS}\right) = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_p X_p$$
$$PS = \frac{\exp(\beta_0 + \beta_1 X_1 + \dots + \beta_p X_p)}{1 + \exp(\beta_0 + \beta_1 X_1 + \dots + \beta_p X_p)}$$

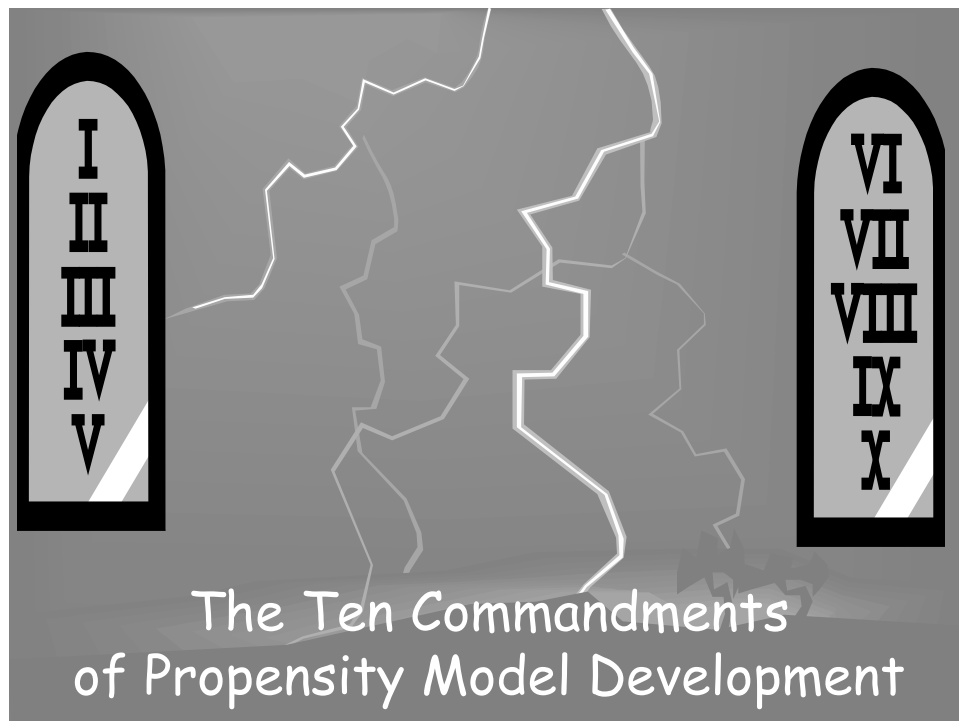
The New Database, Simply

Subject	Propensity to get Treatment	Outcome under Treatment	Outcome under No Treatment
A	.80	12	?
B	.50	7	?
C	.50	?	3
D	.80	?	9

- Match A to D and B to C and plug in resulting estimates for question marks...

So How Do We Build a Propensity Score Model?

- Need to estimate $\Pr(\text{treatment} \mid \text{covariates})$
- Usual tool: Logistic regression model for the treatment allocation decision
 - We therefore want to consider including any variables that have a relationship to the treatment decision (i.e. precede it in time, and are relevant)
 - Clearly, there's no information included on the actual treatment received, or on the outcome(s).



STRATEGIES FOR USING PROPENSITY METHODS WELL

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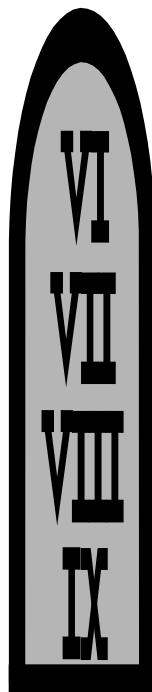
Thou Shalt Value Parsimony

Thou Shalt Examine Thy
Predictors For Collinearity

Thou Shalt Test All Thy Predictors
For Statistical Significance

Thou Shalt Have Ten Times
As Many Subjects As Predictors

Thou Shalt Carefully Examine Thy
Regression Coefficients (Beta Weights)

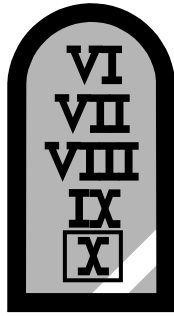
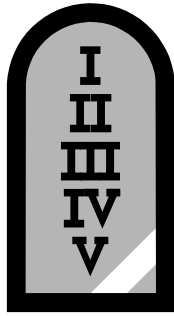


Thou Shalt Perform Bootstrap
Analyses To Assess Shrinkage

Thou Shalt Perform Regression
Diagnostics and Examine Residuals
With Care

Thou Shalt Hold Out A Sample of Thy
Data for Cross-Validation

Thou Shalt Perform External Validation
on a New Sample of Data



Thou Shalt **IGNORE**
Commandments 1 through 9...
And Instead Simply Ensure
That The Model Adequately
Balances The Covariates

Apologies to Joe Schafer

Using the Propensity Score: Three Examples

- Matching using the Propensity Score
 - Aspirin Use and Mortality in Stress Echo patients
- Stratification / Subclassification using the Propensity Score (also includes weighting)
 - Right Heart Catheterization in the Seriously Ill
- Direct Adjustment with Propensity Scores
 - Prostate Cancer surgery vs. Radiotherapy

Using the Propensity Score: Multivariate Matching

- Match subjects so that they balance on multiple covariates using one scalar score.
- Goal: Emulate a RCT in matching, then use standard analyses to compare matched sets.
- Design: Treated subjects matched to people who didn't receive treatment but who had similar propensity to receive treatment (match the treated to untreated "clones").

Seminal paper: Rosenbaum and Rubin (1985)

Aspirin Use and Mortality

- 6174 consecutive adults at CCF undergoing stress echocardiography to evaluate coronary disease.
 - 2310 (37%) were taking aspirin (treatment).
 - 31 covariates – demographics, clinical history, medications, cardiovascular assessment, exercise capacity
- Outcome: All-cause mortality (median follow-up: 3.1 y)
 - Univariate Analysis: 4.5% of aspirin patients died, and 4.5% of non-aspirin patients died...
 - Unadjusted Hazard Ratio: 1.08 (0.85, 1.39)
- Logistic regression for Propensity for aspirin use: 31 covariates (good discrimination; $C = 0.83$)

Gum et al. (2001)

How Were The Aspirin Subjects Matched?

- Tried to match each aspirin user to a unique non-user with a PS identical to 5 digits.
- If not possible, proceeded to a 4-digit match, then 3-digit, 2-digit, and finally a 1-digit match (i.e., propensity scores within .099).
- Result: matches for 1351 (58%) of the 2310 aspirin patients to 1351 unique non-users.

SAS macro: <http://www2.sas.com/proceedings/sugi26/p214-26.pdf>

Baseline Characteristics By Aspirin Use (in %) (before matching)

Variable	Aspirin (n = 2310)	No Aspirin (n = 3864)	P value
Men	77.0	56.1	< .001
Clinical history: diabetes	16.8	11.2	< .001
hypertension	53.0	40.6	< .001
prior coronary artery disease	69.7	20.1	< .001
congestive heart failure	5.5	4.6	.12
Medication use: Beta-blocker	35.1	14.2	< .001
ACE inhibitor	13.0	11.4	< .001

- Baseline characteristics appear very dissimilar: 25 of 31 covariates have $p < .001$, 28 of 31 have $p < .05$.
- Aspirin user covariates indicate higher mortality risk.

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Baseline Characteristics By Aspirin Use [%] (after matching)

Variable	Aspirin (n = 1351)	No Aspirin (n = 1351)	P value
Men	70.4	72.1	.33
Clinical history: diabetes	15.0	15.3	.83
hypertension	50.3	51.7	.46
prior coronary artery disease	48.3	48.8	.79
congestive heart failure	5.8	6.6	.43
Medication use: Beta-blocker	26.1	26.5	.79
ACE inhibitor	15.5	15.8	.79

- Baseline characteristics similar in matched users and non-users.
- 30 of 31 covariates show NS difference between matched users and non-users. [Peak exercise capacity for men is p = .01]

Using Standardized Differences to Measure Covariate Balance

$$d = \frac{100(\bar{x}_{Treatment} - \bar{x}_{Control})}{\sqrt{\frac{s_{Treatment}^2 + s_{Control}^2}{2}}}$$

for continuous variables

$$d = \frac{100(p_{Treatment} - p_{Control})}{\sqrt{\frac{p_T(1-p_T) + p_C(1-p_C)}{2}}}$$

for binary variables

Absolute Standardized Differences > 10%

indicate serious imbalance [Normand et al. (2001)]

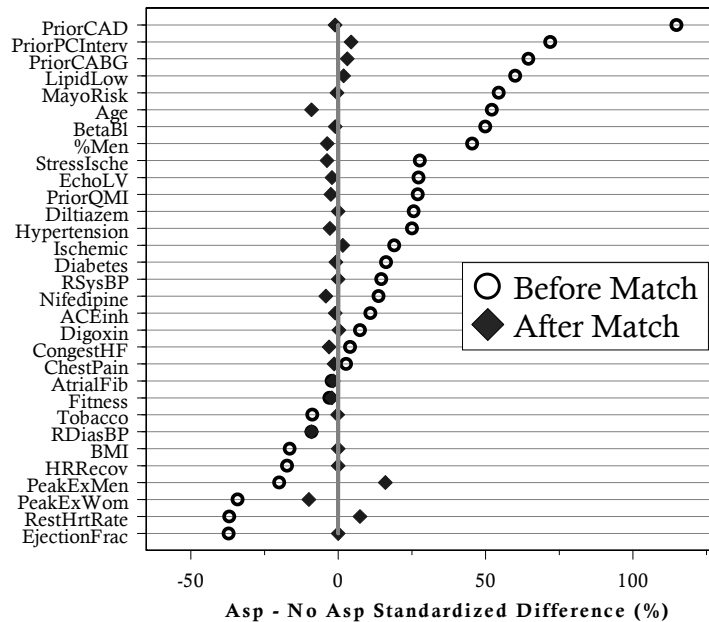
	Aspirin	No Aspirin	P	Std. D.
Before Match	35.1% [811/2310]	14.2% [550/3864]	< .001	49.9%
After Match	26.1% [352/1351]	26.5% [358/1351]	0.79	-1.0%

STRATEGIES FOR USING PROPENSITY METHODS WELL

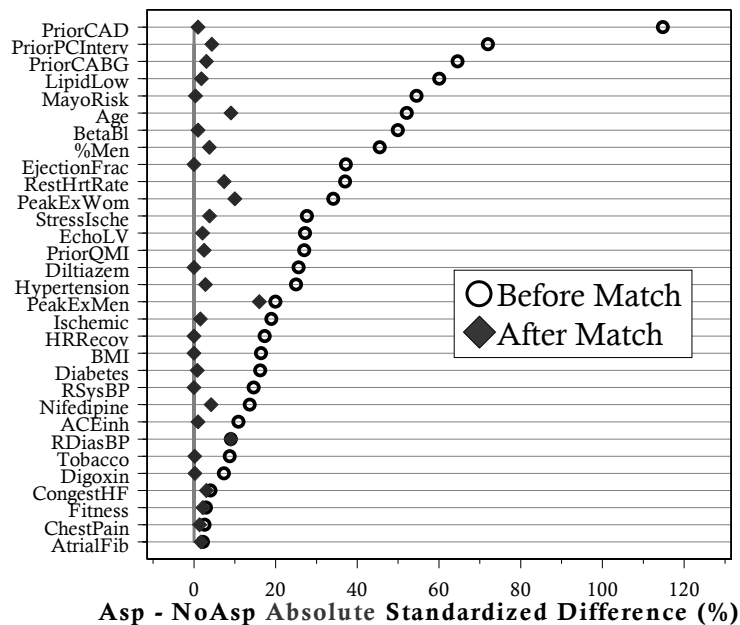
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Covariate Balance for Aspirin Study



Absolute Standardized Differences



Matching with Propensity Scores

- 1351 aspirin subjects matched well to non-aspirin subjects – big improvement in covariate balance. Matched group looks like an RCT...
- Matching still incomplete, but results on PS matched group mirrored the results for the covariate-adjusted group as a whole...
- Resulting matched pairs analyzed using standard statistical methods, e.g. Kaplan-Meier, Cox proportional hazards models.

Estimating The Hazard Ratios

Approach	n	Hazard Ratio	95% CI
Full sample, no adjustment	6174	1.08	(.85, 1.39)
Full sample with no PS, adjusted for all covariates	6174	0.67	(.51, .87)
PS-Matched sample	2702	0.53	(.38, .74)
PS-Matched, adjusted for PS and all covariates	2702	0.56	(.40, .78)

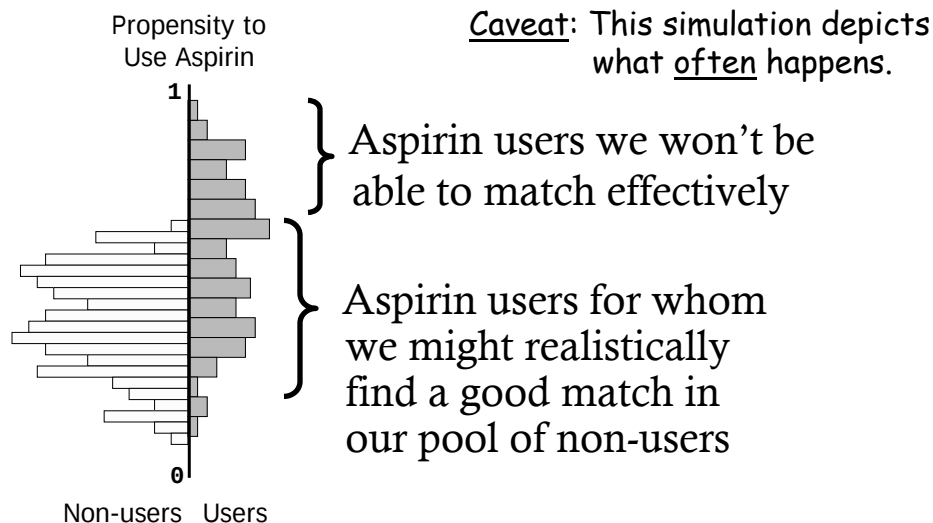
- During follow-up 153 (6%) of the 2702 propensity score-matched patients died.
- Aspirin use was associated with a lower risk of death in matched group (4% vs. 8%, $p = .002$).

Incomplete vs. Inexact Matching

- Trade-off between
 - Failing to match all treated subjects (incomplete)
 - Matching dissimilar subjects (inexact matching)
- Severe bias due to incomplete matching – it's usually better to match all treated subjects, then follow with analytical adjustments for residual imbalances in the covariates.
- In practice, concern has been inexactness.
- Certainly worthwhile to define the comparison group and carefully explore why subjects match.

For more, see Rosenbaum (1985, 2002)

Who's Getting Matched Here? Where Do The Propensity Scores Overlap?



Using the Propensity Score to Stratify (Subclassify) Subjects

- Stratification by Propensity Score Quintile
 - Fit a PS model for each subject
 - Split the subjects into 5 strata (subclasses) of equal size by their propensity scores.
- Five strata of equal size (quintiles) constructed from the PS will usually suffice to remove over 90% of the selection bias due to each of the individual covariates in the PS model.

Seminal paper: Rosenbaum and Rubin (1984)

Right Heart Catheterization and Mortality

- Why an OS? RHC very popular – Equipoise?
 - RHC directly measures cardiac function – lots of reasons to think this would be helpful.
 - Physician makes the decision – ethical to participate?
 - RHC effect sizes may be small
- 5735 seriously ill hospitalized pts in SUPPORT
 - 2184 treated patients (RHC within 24 h of admission)
 - 3551 controls (no RHC in first 24 h after admission)
- Key Outcome: 30 day survival

Connors et al. (1996)

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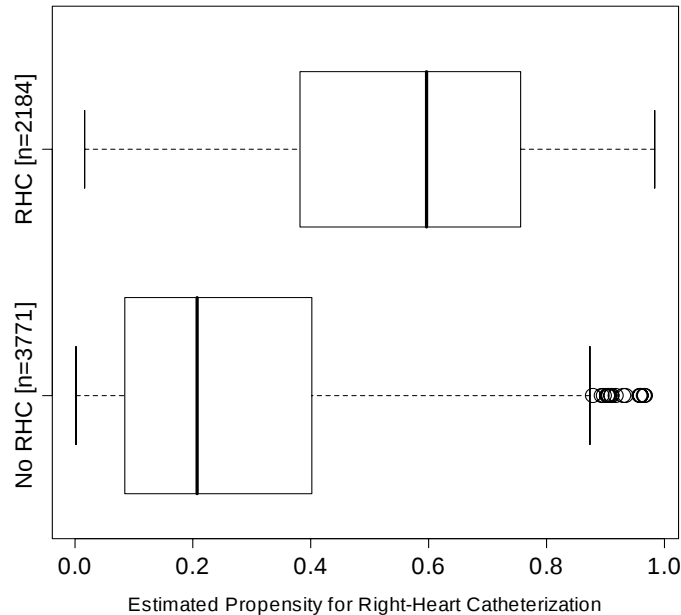
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Some Characteristics used to predict Propensity for RHC use

- Age, Sex, Race
- Education, Income
- Insurance type
- Primary and secondary disease category
- 12 categories of admission diagnosis
- ADL & DASI 2 weeks before admission
- DNR status on day 1
- Cancer (none, local, metastatized)
- 2-month survival model
- Weight, temperature, BP, heart rate, resp. rate
- 13 categories of comorbid illness
- Body chemistry (pH, WBC, PaCO₂, etc.)

Panel (7 specialists in clinical care) specified important variables related to the decision to use or not use a RHC.

Boxplot of Propensity Score Overlap

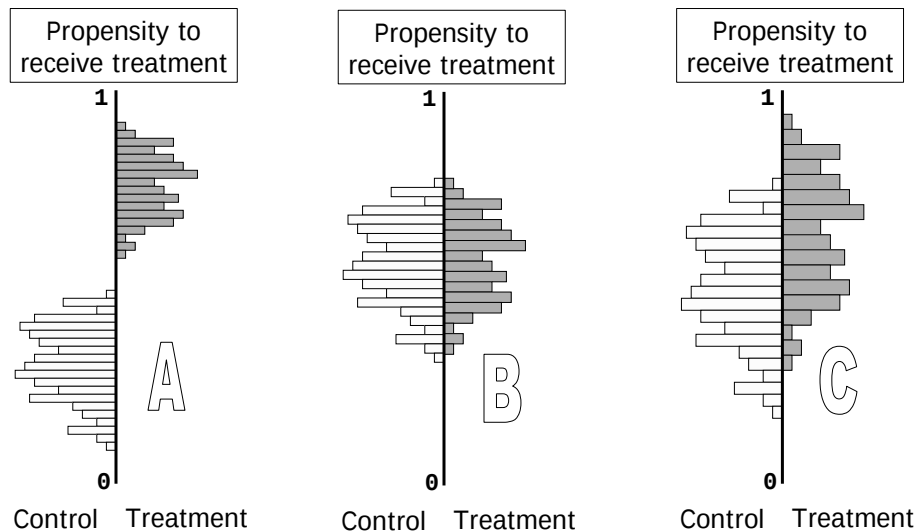


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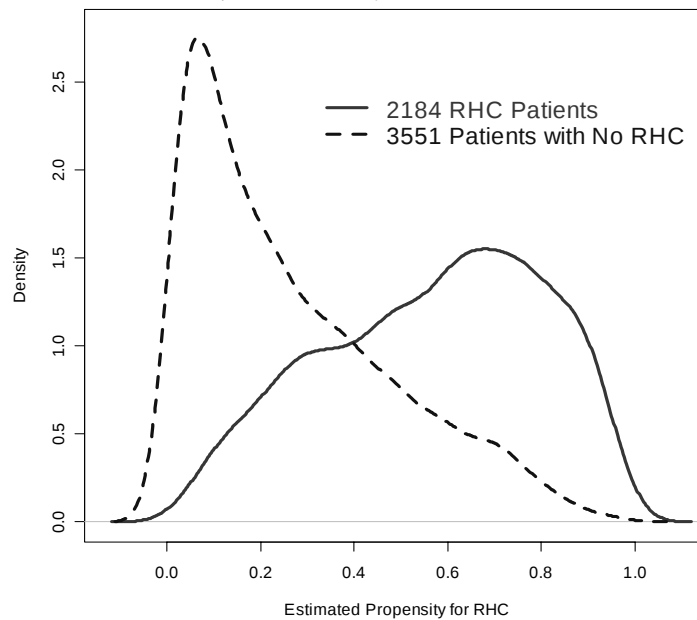
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How Much Propensity Score Overlap Do We Want?



Do the Propensity Scores Overlap?

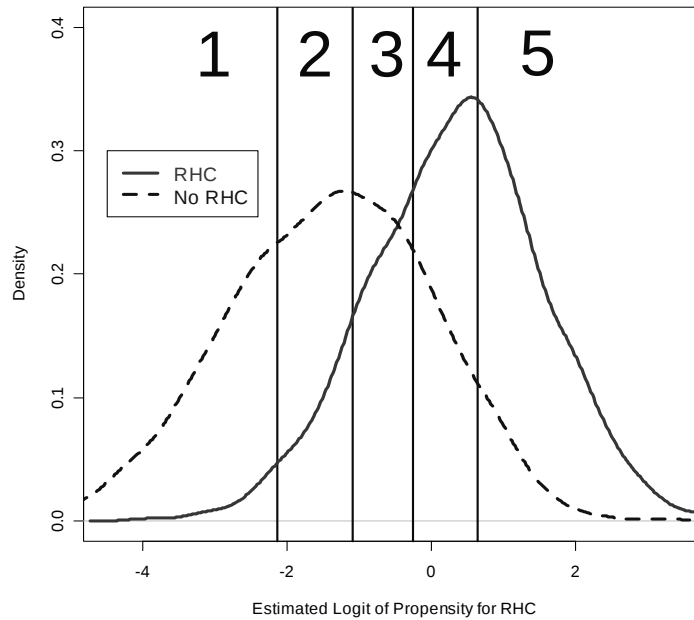


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Subclassification by PS Quintiles



Propensity Score Quintile Subclassification

- 5735 patients were divided into five strata of 1147 patients each, by estimated linear propensity scores.

PS Subclass	Propensity Score $\approx$ Prob(RHC covars.)	Actually got RHC	Actually got no RHC
5	Highest 1147 scores	888 (77%)	259 (23%)
4	2 nd highest	635 (55%)	512 (45%)
3	Middle	398 (35%)	749 (65%)
2	2 nd lowest	210 (18%)	937 (82%)
1	Lowest 1147 scores	53 (5%)	1094 (95%)

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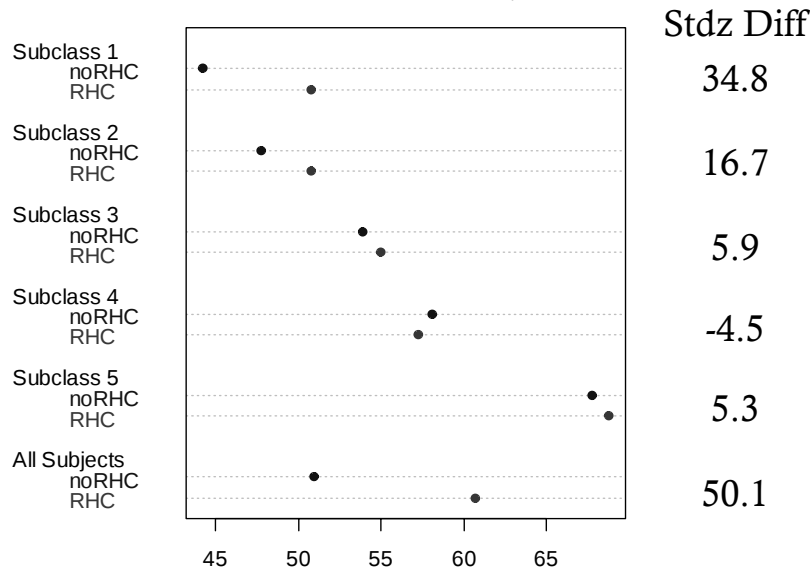
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Subclassification & Covariate Balance

Age>80	Overall	S1	S2	S3	S4	S5
RHC (%)	7.6	18.9	11.0	10.8	7.2	5.1
noRHC (%)	14.1	22.2	14.1	11.1	5.9	4.6
Stzd. Diff.	-20.8	-8.3	-9.5	-0.9	5.6	2.0
p-value	< .00001	0.73	0.26	0.92	0.40	0.87

Mean DBP	Overall	S1	S2	S3	S4	S5
RHC	68	94	86	78	70	57
noRHC	85	100	83	77	72	59
Stzd. Diff.	-45.5	-13.4	-5.4	3.1	-6.5	-6.2
p-value	< .00001	0.33	0.47	0.62	0.28	0.37

APACHE III Acute Physiology Score Across Subclasses



Subclass Specific Estimates

PS Subclass	Group	Patients	P(Survive 30 d)
1	RHC	53	.698
	No RHC	1094	.705
2	RHC	210	.619
	No RHC	937	.705
3	RHC	398	.643
	No RHC	749	.706
4	RHC	635	.649
	No RHC	512	.688
5	RHC	888	.595
	No RHC	259	.645

How Should We Stratify?

- Hullsieck and Louis (2002) suggest balancing iteratively on the inverse variance of the subclass-specific treatment effects, rather than more simply on the number of observations within a subclass.
 - Having lots of subclasses reduces bias due to confounding.
 - Similar variances of treatment effect across subclasses maximizes effective bias control.

Hullsieck and Louis (2002)

Propensity Score Weighting

- Idea: Re-weight treated and control observations to make them representative of the population of interest
- A treated subject's weight is the inverse of its propensity score.

$$w_i = \frac{1}{PS_i}$$

- A control subject's weight is the inverse of 1 minus its propensity score.

$$w_i = \frac{1}{(1 - PS_i)}$$

Rubin (2001), Rosenbaum (1987), Lunceford and Davidian (2004)

Right Heart Catheterization

- 5735 hospitalized patients in the SUPPORT study: 2184 treated (RHC) and 3551 controls (no RHC).
- Reweight each treated patient by 1/PS, and each control patient by 1/(1-PS).
- PS model estimated by Hirano and Imbens using 57 of 72 available covariates (those with $t > 2.0$).

Hirano and Imbens (2001), Connors (1996), Hirano, Imbens and Ridder (2003)

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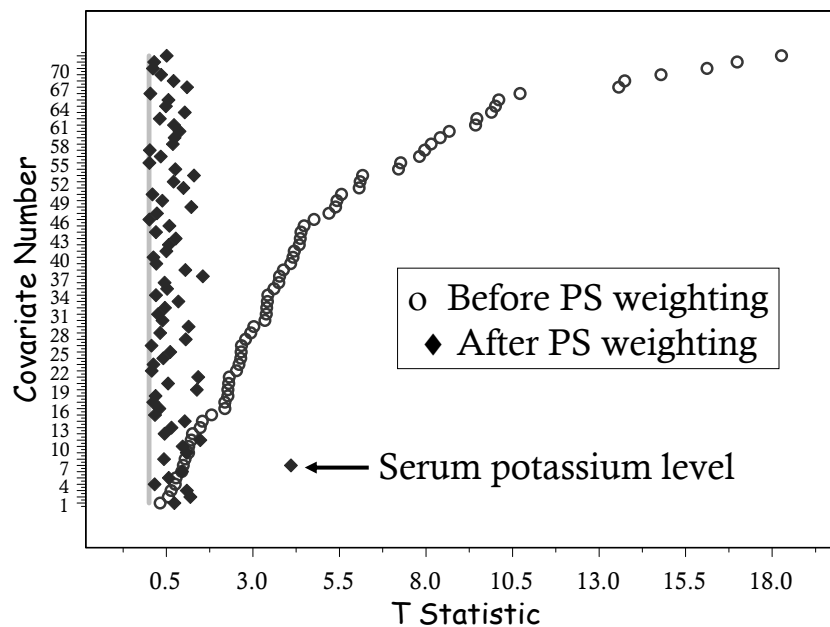
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RHC Covariate Balance Before/After Weighting on the Propensity Score

Variable	Before Weighting			After Weighting		
	No RHC	RHC	t	No RHC	RHC	t
Age	61.76	60.74	-2.28	61.25	61.15	-0.19
Sex	0.46	0.41	-3.42	0.44	0.43	-0.85
Edu	11.56	11.85	3.35	11.68	11.71	0.39
COPD	0.11	0.02	-13.6	0.07	0.06	-1.10
Album	3.16	2.97	-8.15	3.08	3.15	0.69
Potass	4.07	4.04	-0.99	4.15	3.97	-4.10

Hirano and Imbens (2001)

Absolute T Statistics for RHC vs. No RHC Group Means



What To Include in the PS Model

- All covariates that subject matter experts (and subjects) judge important when selecting treatments.
- All covariates that relate to treatment and outcome, certainly including any covariate that improves prediction (of exposure group).
- Sop up as much “signal” as possible.

Estimating the Treatment Effect

- Weighting: attractive theoretical properties, but rarely used (so far) in practice.
- Outcome model (logistic regression) is estimated using MLE with IPS weights
- Z is exposure assignment vector, X is vector of covariates for susceptibility adjustment, weights are as described above
- Estimated treatment effect:

Joffe et al. (2004) or Austin and Mamdani (in press) provide nice applications.

Using the Propensity Score in Direct Covariate Adjustment

- We sometimes use treatment indicator and the propensity score itself as predictors, along with other important covariates and risk adjusters.
- It's common to take a large set of covariates to form the PS, then use the PS and a subset of those covariates in the model for outcomes.
 - Analogous to matching within PS calipers, choosing the best match in terms of distance on key covariates from the target from among subjects inside the calipers.
- We could also regress within PS strata, etc.

Prostate Cancer Surgery vs. Radiotherapy

- RP: Radical prostatectomy (n = 1156)
- RT: external beam radiotherapy (n = 435)
- Outcomes: relevant functioning, QOL
- PS = Pr(RP) model built for 26 covariates: used survey weights, and imputed missing data.
- Split into PS quintiles to check balance – no sig diffs. in covariate means, and reasonable overlap of PS.
- PS used as covariate in cross-sectional and longitudinal regressions comparing RP to RT.

Potosky et al (2000)

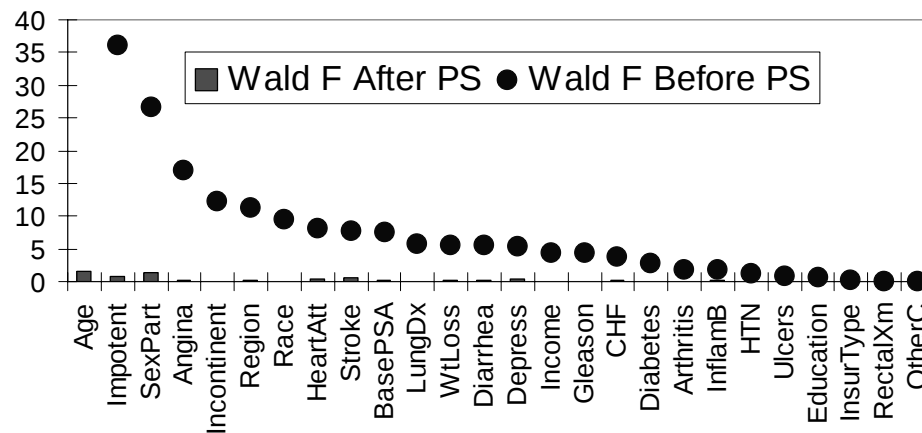
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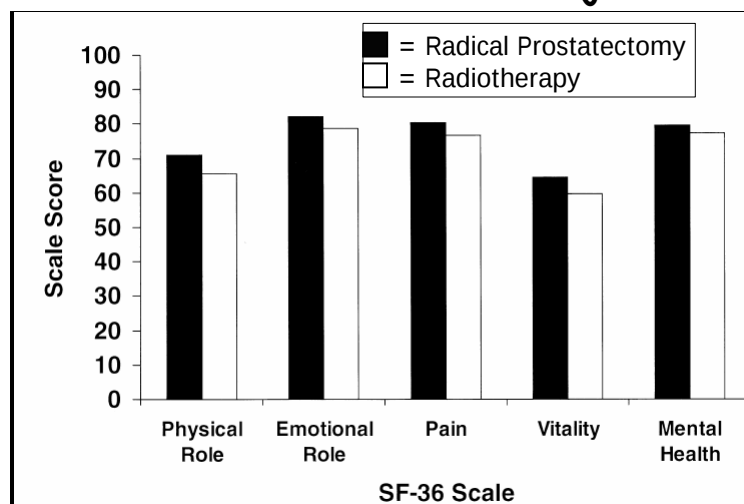
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Are the Covariates Balanced? Before/After PS Stratification



Comparing Outcomes Directly After PS and Covariate Adjustment



No significant differences here.

Could an Unmeasured Covariate Be Responsible for these Conclusions?

- It is unlikely that a hidden bias would substantially affect these conclusions:
 - Measured and incorporated every major known factor that they could identify.
 - Treatment effects on health outcomes were generally quite large, consistent with earlier studies, and clinically plausible.
- LTE: Logic, Theory & Empirical Evidence

Conclusions of Sensitivity Analysis in terms of an unobserved covariate

- When describing possible hidden bias, we refer to characteristics we did not observe, and therefore did not control for in the PS model.
- If our study was randomized, or somehow free of hidden bias, we would have strong evidence of a treatment effect.
- Structure of Argument: To explain away the observed effect, an unobserved covariate would need to increase the odds of exposure to treatment by more than a factor of ____.

Rosenbaum (1991, 2002), Rosenbaum and Rubin (1983)

Summary: Sensitivity Analysis

- Hidden bias is the great problem with observational studies, and for PS models.
- Sensitivity analysis can be applied to many statistical tests – we hope to find that an unobserved covariate would have to be very powerful to alter conclusions.
- This doesn't mean that such a covariate (or set of them) doesn't exist.

What Propensity Scores Can and Cannot Do

- If we match (subclassify) treated subjects to controls with similar propensity scores, we can behave as if they had been randomly assigned to treatments.
- Or, if we use regression to adjust for propensity to get treatment, we can compare treated to controls without worrying about the impact of baseline differences we've measured on selection to exposure.
- But if our propensity model misses an important reason why subjects are selected to an exposure, we'll be in trouble, and we'll never know it.

Part Two

Estimating the Propensity Score Effectively

Building the Propensity Score Model

- If propensity scores are to be used in design to balance covariates, it is the distributional balance of those covariates that is achieved within blocks (strata, subclasses or matched pairs) that is the critical diagnostic tool.

Donald B. Rubin

Rubin (2004)

Should Adjustments Be Made for All Observed Covariates?

- If not, how should covariates be selected?
- No real reason to avoid adjustment for a true covariate – a variable describing subjects before treatment.
- In practice, though, this can increase both cost and complexity unnecessarily.
- There are issues of data quality and completeness to consider.

DON'T Select Covariates Like This...

Covariate	Treatment	Control	T test Sig?	Include?
Age	49.8	50.1	No, $p > .05$	No
BMI	25.5	23.2	Yes	Yes
Heart Rate	82	76	Yes	Yes
Years Ed.	11.3	12.0	No, $p > .05$	No
Systolic BP	135	133	No, $p > .05$	No
SF-36 Phys	61	53	Yes	Yes

- Most common method: Compare treated and control groups on a long list of covariates, with a t test. Then adjust only for those with a significant difference.

Why Not Screen This Way?

- Doesn't consider relationship between covariate and outcome.
- No reason to believe that the absence of statistical significance implies the imbalance of the covariate is small enough to be ignored.
- Process considers covariates one at a time, while the adjustments will control the covariates simultaneously.

OK, What Should We Do?

- Give the data multiple opportunities to call attention to potential problems.
- Select a tentative list of covariates for adjustments using problem knowledge and exploratory comparisons of treatment groups.
- Select tentative adjustment method and apply it to the covariates excluded from the list, identifying large imbalances after adjustment.
- Reconsider the tentative list in light of this.

Review of 47 Published Articles

- **Variable selection in the PS model**
 - 24 provided no information about variable selection
 - 6 used non-parsimonious logistic regression models
 - 7 identified variables through univariate sig. testing
 - 5 studies used stepwise or other selection algorithms
 - 4 studies used a priori (not data driven) selection
 - 1 study selected variables based on GOF tests
- **Interaction Inclusion Criteria**
 - 30 of 47 left use of interactions in PS model unclear
 - 12 clearly indicated that no interactions were used
 - 3 of the other 5 used p values to select interactions
 - 1 used improvement in discrimination of the PS model
 - 1 used improvement in balance across exposure groups

Weitzen et al (2004)

What People Have Done

- **Adequacy of the Propensity Score Model**
 - Goodness of Fit
 - 6 of the 47 studies considered the GoF of the PS model
 - 4 of the 6 provided the Hosmer-Lemeshow test p value
 - When the H-L was NS, one group didn't use the PS
 - 1 study used split-sample validation to evaluate fit
 - 1 study indicated model was “appropriately calibrated”
 - Discrimination
 - 18 of 47 studies reported C statistics (range: 0.52, 0.92)
 - Balance on Covariates
 - 22 of 47 included no information on covariate balance
 - In one article, when there was no balance in 3 of the 5 PS quintiles, they excluded patients from those quintiles

Why not model outcome using all variables in the propensity model?

- Two stages: fit PS, then use PS in model
- One stage: just fit big outcome model
- Pros of two-stage approach:
 - Forces you to think hard about selection.
 - You don't care about parsimony in the PS, so you get maximum predictive value there.
 - You can fit a very complicated PS model first with interactions, higher order terms, splines, etc.
 - You can fit a smaller outcome model, which may let you assess its validity more accurately.

What about Propensity Model Diagnostics?

- Rubin describes “confusion between two kinds of statistical diagnostics...”
 - (1) Diagnostics for the successful prediction of probabilities and parameter estimates underlying those probabilities
 - (2) Diagnostics for the successful design of observational studies based on estimated propensity scores.
- Basically, (2) has a role – (1) doesn't, here.

Rubin (2004)

Should we be checking propensity model goodness of fit?

- Are tests used to evaluate logistic model fit and discrimination helpful in detecting the omission of an important confounder?
 - Simulated data including an important binary confounder – compared inclusion to exclusion
- Hosmer-Lemeshow GOF and C statistic were of no value in detecting residual confounding in treatment effect estimates

Weitzen et al. (2005)

Dealing with Hidden Bias?

- Assess potential for unmeasured covariates to affect results with sensitivity analysis.
- The unmeasured covariate in question would have to be independent of all variables in the propensity model.
 - Either we missed a domain of the problem
 - Or our measure is so weak that we miss something important completely.

Does Propensity Matching Balance "Omitted" Covariates?

- We fit a published propensity model to data from the SUPPORT study on right heart catheterization, which used 82 covariates.
- Then we got data on 17 other covariates, not included in the propensity model.

Correlation with Propensity Score?	# of Covariates	Balance Improved After Match	Median Bias Reduction
Significant ($\alpha = .05$)	10	9 (90%)	45%
Not Significant	7	2 (29%)	-36%

Love, Cebul, Thomas and Dawson (2003), Connors et al. (1996)

Conclusions: When is an omitted variable most dangerous?

- All observational studies are potentially affected by hidden bias. Sensitivity analyses are necessary in any such study.
- An omitted variable is most likely to affect conclusions about the exposure if it is:
 - closely related to outcome.
 - seriously imbalanced by exposure.
 - uncorrelated with propensity score.

Missing Data and Generalized Propensity Scores

- Pattern of missing covariates can be prognostically important
- PS should, ideally, condition both on
 - Observed values of covariates, and
 - Observed missing-data indicators
- Generalized PS lead in expectation to balanced missing data patterns & covariate distributions in the treatment and control groups

D'Agostino and Rubin (2000), D'Agostino and Lang (2001)

Dealing with Missing Data: A Typical Approach

- MAR = assumed missing at random – mechanism by which data are missing is unrelated to information not in **X**.
 - Discrete: include binary “missing” x
 - Continuous: fit two predictors:
 - [1] subject was measured/unmeasured,
 - [2] if subject measured, then value.

Part Three

Assessing and Displaying Covariate Balance

Checking for Covariate Balance

Nursing Home Study

- Rx: Rehabilitation
- Control: No Rehab
- 21 covariates included in PS model (no interactions or polynomial terms)
- Matched patients with similar propensities for rehabilitation

Murray et al. (2003, 2005)

Variable	Rehab	No Rehab
Mean Age	78.1	78.0
% female	67	64
Mean BMI	24.2	23.6
Mean Cog ADL	79.6	71.0
% ath. ♥ disease	15	16
% dementia	20	38
% live alone	40	33
% with a cane	54	30
% recent ADL ↓↓	66	41

STRATEGIES FOR USING PROPENSITY METHODS WELL

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Checking for Covariate Balance

Variable	Rehab	No Rehab	Matched Rehab	Matched No Rehab
Mean Age	78.1	78.0	78.1	78.0
% female	67	64	65	65
Mean BMI	24.2	23.6	23.6	23.6
Mean Cog ADL	79.6	71.0	75.0	74.6
% ath. ♥ disease	15	16	18	17
% dementia	20	38	28	29
% live alone	40	33	37	36
% with a cane	54	30	37	38
% recent ADL ↓↓	66	41	52	52

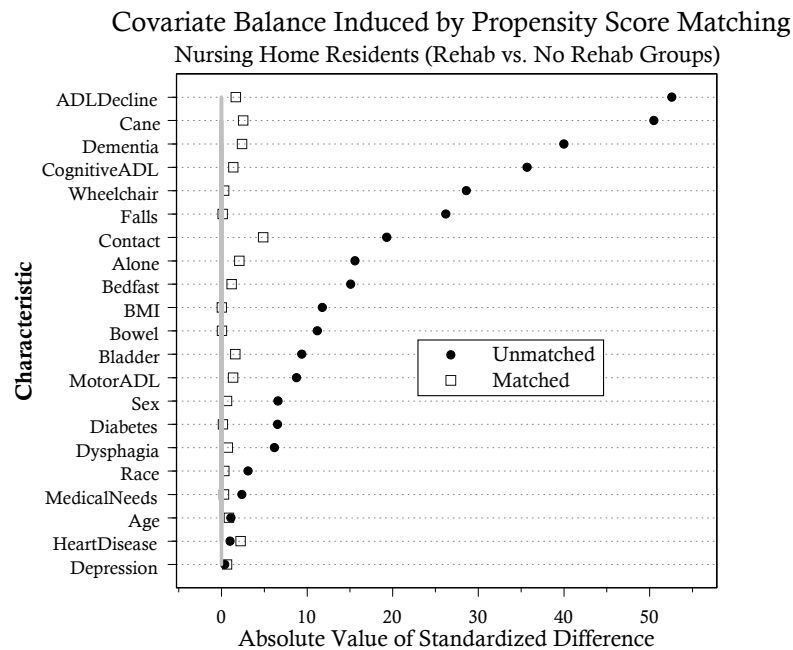
Choices To Be Made

- How should I summarize balance within a covariate?
 - What if I have both continuous and categorical covariates?
 - Standardized differences
 - Significance test results
- How can I best present the summarized results across covariates?
- What are the key messages to get across?

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Using Propensity Scores in Designing Observational Studies

- Specify treatments, outcomes, covariates.
- Collect treatment and covariate information, and model treatment assignment with the propensity score.
- Use propensity scores (through matching, stratification, reweighting) to reduce bias.
- Check for covariate balance across the treatment groups and iterate until happy...

Rubin (1997, 2001)

Rich and Poor Covariate Sets

- With a rich set of covariates, adjustments for hidden covariates may be less critical.
- With less rich covariate sets, we may need to do more – say, try to find an instrument.
- Conclusion after the initial design stage may be that the treatment and control groups are too far apart to produce reliable effect estimates without heroic modeling assumptions.

When Can We Move On?

- Three conditions which must all apply for regression adjustment to be trustworthy:
- Difference in the means of $\text{logit}(\text{PS})$ in the two groups being compared must be small.
- Ratio of variances of $\text{logit}(\text{PS})$ in the two groups must be close to 1.
- Ratio of variances of the “residuals” of the covariates after PS adjustment close to 1.

Why Work This Hard in the Initial Stage of Design?

- No harm, no foul.
 - Since no outcome data are available to the PS, nothing based on the PS biases estimation of treatment effects.
- Balancing covariates / PS makes subsequent model-based adjustments more reliable.
 - Model adjustments can be extremely unreliable when treatment groups are far apart on covariates.

Comparing “Current Smokers” to “Never Smokers” using the National Medical Examination Survey

- Large nationally representative data base of nearly 30,000 adults, calendar year 1987
- Design Goal: Create samples of smokers and never smokers in NMES with the same multivariate distribution of covariates to assess causal effects
- Genders separately assessed. We focus on Male “Current Smokers” vs. Male “Never Smokers”
 - 3510 Male “Current Smokers” in the pool
 - 4297 Male “Never Smokers” as controls

Rubin (2001)

Propensity Model for “Current Smokers” vs. “Never Smokers”

- Logistic Regression with sampling weights
- 146 variables in the model including
 - Main effects
 - Quadratic effects
 - Interaction effects
- Separate models were built for “former” vs. “never” and for the female comparisons

Assessing the Degree of Overlap: Step 1: Looking for Mean Bias

- Bias B = standardized difference in the means of logit(propensity scores) between current smokers and never smokers for males
- We want the bias in the propensity score to be small, no greater than 0.50 in absolute value.
- Here, mean propensity score Bias B = 1.09
- In fact standardized difference > 0.5 for many of the individual covariates, as well.

Assessing the Degree of Overlap: Step 2: Looking at the Variance Ratio

- Ratio R = ratio of the variances of $\text{logit}(\text{propensity scores})$ between current smokers and never smokers for males.
- We want the variances to be homogeneous, so the ratio should be close to 1 ($1/2$ and 2 are far too extreme).
- Here, variance ratio for $\text{logit}(PS)$ is $R = 1.00$
- Could look at ratio of covariate variances, also

Assessing the Degree of Overlap: Step 3: Comparing Residuals

- Regress each covariate on $\text{logit}(PS)$ and look at ratio of variances of residuals for current smokers to variance of residuals for never smokers within the male population.
- Here, we get a separate result for each of the 146 covariates – we want results near 1.00
- 57% of covariates had $0.8 \leq \text{resid ratio} \leq 1.2$
- 5% of covariates had $\text{resid ratio} > 2$ or < 0.5

Mahalanobis Matching of Smokers & Never Smokers within PS Calipers

- For the 3510 male current smokers, 3510 “matching” male never smokers were chosen from the pool of 4297 male never smokers.
- Method: Mahalanobis metric matching within propensity score calipers (± 0.2 of a linear propensity score standard deviation).
- The four Mahalanobis distance variables used to do the matching were: age, education, body mass index, and sampling weight.

Impact of Matching on Overlap: Current vs. Never Smoking Males

- Before The Match

		Residual Variance Ratios					Total
B	R	$\leq .5$	.5 - .8	.8 - 1.2	1.2 - 2	> 2	
1.09	1.00	3	9	57	26	5	100

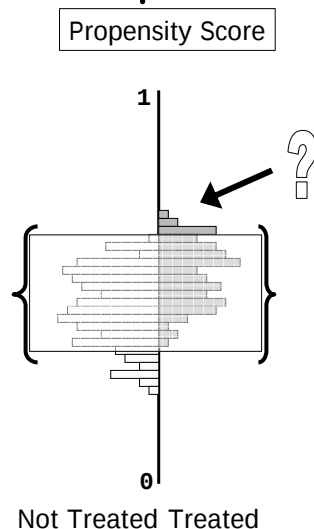
- After The Match

B	R	$\leq .5$	.5 - .8	.8 - 1.2	1.2 - 2	> 2	Total
0.08	1.16	1	3	90	6	0	100

What Do We Do If Some Treated Subjects Can't Be Matched Well?

- In this case, there were no “current smoker” Males that could not be matched within the PS calipers to “never smoker” Males.
- What if there had been a treated subject whose propensity score was not “matchable?”
- What if the “donor pool” of never smokers had been empty for one of the current smokers?

What if Treated and Untreated Groups Don't Overlap Completely?



- Inferences for the causal effects of treatment on the subjects with no overlap cannot be drawn without heroic modeling assumptions.
- Usually, we'd exclude these treated subjects, and explain separately.

Subclassification of Matched Samples of Smokers and Never Smokers

- Suppose we still weren't satisfied.
- Idea: Create two equal-size (weighted) subclasses, low and high on logit(PS).
- Treated and Control subjects with low PS are to be compared to each other.
- Treated and Control subjects with high PS are to be compared to each other.
- Weighted average of two comparisons = result.
- No reason to stop at just 2 subclasses, either...

Impact of Post-Matching Subclassification on Overlap

Current vs. never smoking males

			Residual Variance Ratios				
Subclasses	B	R	≤.5	.5 - .8	.8 - 1.2	1.2 - 2	>2
1	0.39	1.33	0	4	88	8	0
2	0.18	1.36	0	2	98	0	0
4	0.10	1.25	0	1	99	0	0
6	0.09	1.30	0	0	100	0	0
8	0.08	1.16	0	0	100	0	0
10	0.07	1.12	0	0	100	0	0

Why Can We Get Away With This?

- We can obtain dramatic reduction in initial bias through these propensity-based methods
- If substantial balance in covariates is obtained in this initial design stage, the exact form of the modeling adjustment is not critical.
- Similar treated and control covariate distributions implies only limited model-based sensitivity.
- Why doesn't this introduce a bias for our eventual conclusions and analytic results?
- Because we haven't got the outcomes.

What Should You Do About Residual Covariate Imbalance?

- Suppose a covariate appears seriously imbalanced after propensity matching.
- Could make a regression adjustment for that covariate after matching.
- Could use an additional or alternative measurement of the concept described by the covariate in the PS model.
- Consider re-matching starting with a different random order of treated patients, or by a different standard.
 - Consider Mahalanobis distance matching within propensity score calipers.

Part Four

Choosing Analytic Techniques

Advice on Getting A Better Match

- Two concerns: Covariate balance, and Matching as many subjects as possible.
 - If match is incomplete, consider both matching and non-matching (stratification/regression) analyses.
 - Match logit(PS) instead of raw PS, accept matches within a fraction (usually .6) of the pooled (across treatments) standard error of the PS. This is more defensible statistically, and should improve yield.
 - Matching on multivariate distance within PS calipers usually beats matching just on PS.
 - Can do OPTIMAL full matching (see Bergstralh et al. macro at <http://www.mayo.edu/hsr/sasmac.html>)

Is Regression Adjustment for Observational Studies Obsolete?

- Matching and stratification are old and trusted methods of adjustment for observational studies, but the difficulty of implementing them led earlier practitioners to prefer regression.
- It is now possible to perform optimal full matching, and to achieve levels of bias reduction that were previously unattainable.

Hansen (2004)

Full Matching in a study of coaching for the SAT

- It has been tough to implement full matching in observational studies, even though it is the best in principle:
- Alignment of comparable treated and control subjects is as good as any alternate method, and potentially much better
- Hansen modifies full matching with modifications to minimize variance as well as bias
 - Optimal full matching removes as much as 99% of the bias along a PS on which treated and control means are separated by 1.1 SD's.
 - Reduces to insignificance biases along **27** covariates, while making use of more, not less, of the data than regression based analyses.

Hansen (2004)

SAT Coaching Study

- Survey of a random sample of 1995-1996 SAT test takers about their preparation
- 12% of respondents had completed extracurricular test preparation courses
- Matching looked unattractive to the original researchers due to significant reduction in sample size – but they only considered 1:1 matching.
- Do 1:k matching options look better?

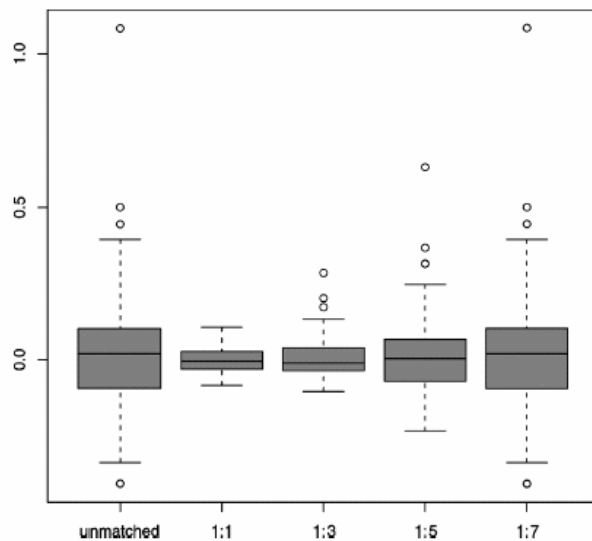


Figure 1. Covariate Imbalances in 1:k Matching. Each boxplot represents standardized biases in the 99 categories of the 27 categorical covariates along with standardized bias in the propensity score (which in each plot is the uppermost outlier). Strictly speaking, the matching represented at far right is not a 1:7 matching but a blend of six 1:6 and 494 1:7 matched sets.

Covariate Imbalances in 1:k Matching

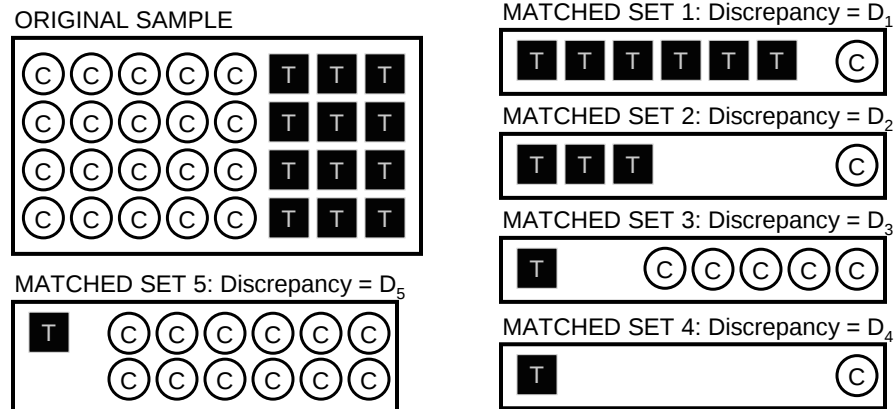
- In all of these cases, we're using less data
- Still some imbalance

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Optimal Full Matching



- OFM minimizes propensity score distances (discrepancies) while using all treated and all control subjects (i.e. discarding no units).
- Here, infinite distances force matches on Race×Sex

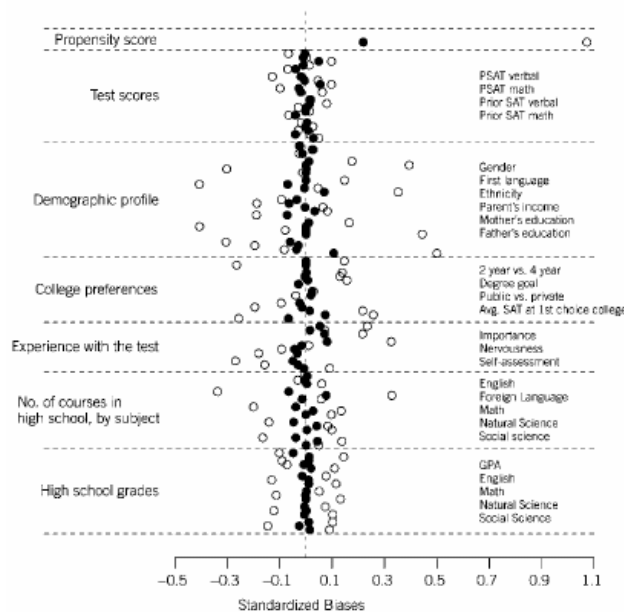


Figure 3. Standardized Biases Without Stratification or Matching. Open Circles, and Under the Optimal [5, 2] Full Match, Shaded Circles.

Standardized Bias Plot

- Open circles are for standardized biases before matching
- Shaded circles describe results after full match

SAT Coaching Study Results

- Raw differences of treated and control group means were 41 points on Math and 9 on Verbal
- Full matching leads to aggregate contrasts of 26 points on Math and 1 point on the verbal.
 - Standard errors for these estimates are around 5 points.
- Surprised that Verbal effect is so small? Recall control condition is not “no prep at all” Estimated effect of treatment on the controls is 3 for Math and -8 on Verbal.
- Method doesn’t require homogeneity of coaching effects – whether and to what degree coaching is beneficial appears to vary greatly across students

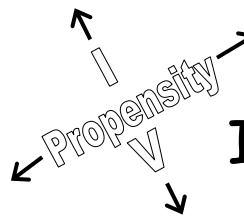
What About Instrumental Variables?

- Idea: Find a variable (the instrument)
 - strongly correlated with the treatment choice
 - but having no direct effect on the outcome (outside of the instrument’s influence on treatment selection)
- If these two conditions are not met, then IV is not a useful approach.
- In health care, treatment selection is usually closely linked to outcome.

See Earle et al (2001), Landrum and Ayanian (2001), Posner et al (2001)

When Are Instrumental Variables Methods Especially Attractive? An instrument is available, and ...

- Assignment to a treatment is ignorable, but compliance with the assignment is not perfect so that the dose of treatment received is non-ignorable.
- Data are weak, in the sense that observed covariates provide insufficient insight into the background to allow estimated effects (adjusting for covariates) to be due to treatment.



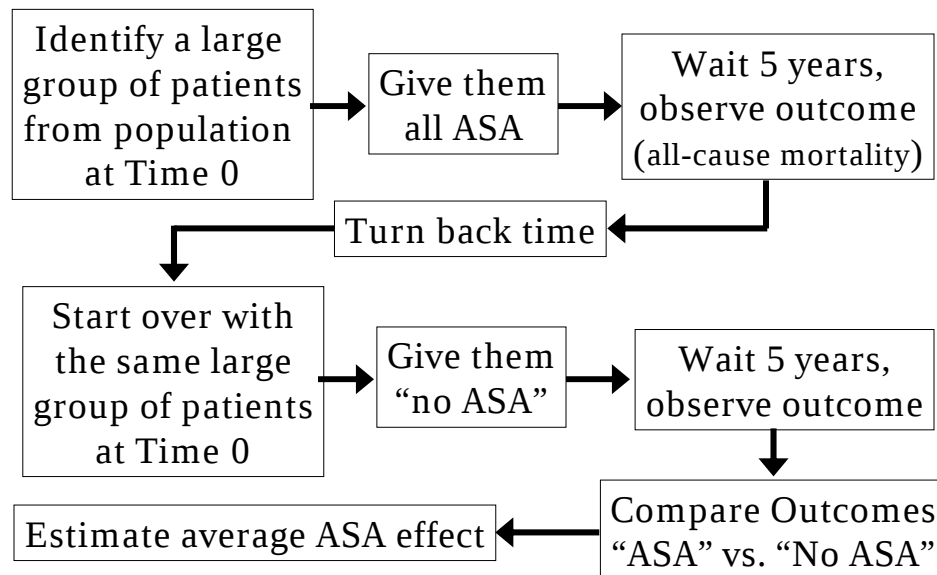
Propensity Scores vs. Instrumental Variables?

- Some questions call for PS adjustment, others for IV models of Rx effect.
- Both have unverifiable assumptions:
 - PS adjusts for selection bias in terms of identified covariates – we must presume this is sufficient to also adjust for unobserved covariates. Sensitivity analysis can help.
 - IV presumes we can and do identify appropriate instrument(s).

Part Five

Communicating Results to a Non-Statistical Audience

ASA and Mortality: Ideal Study

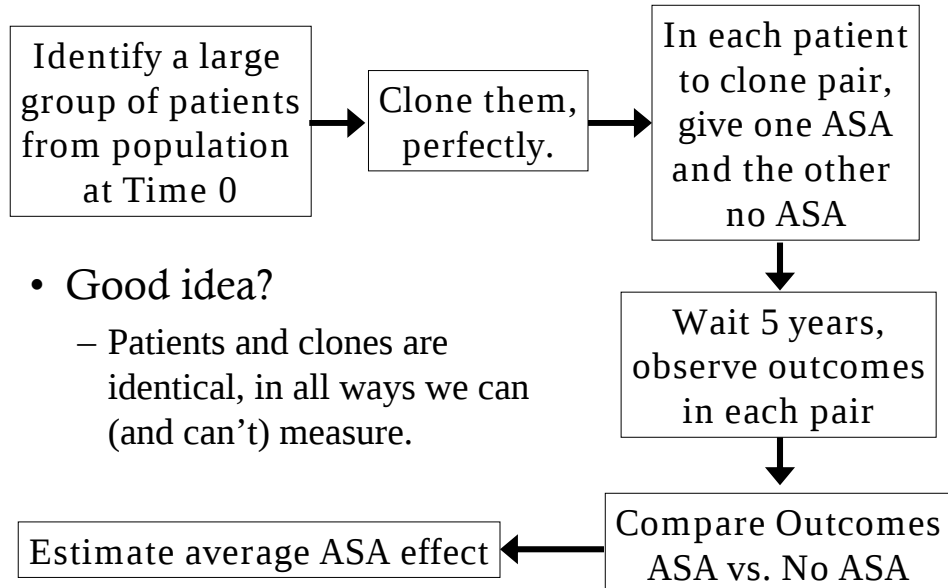


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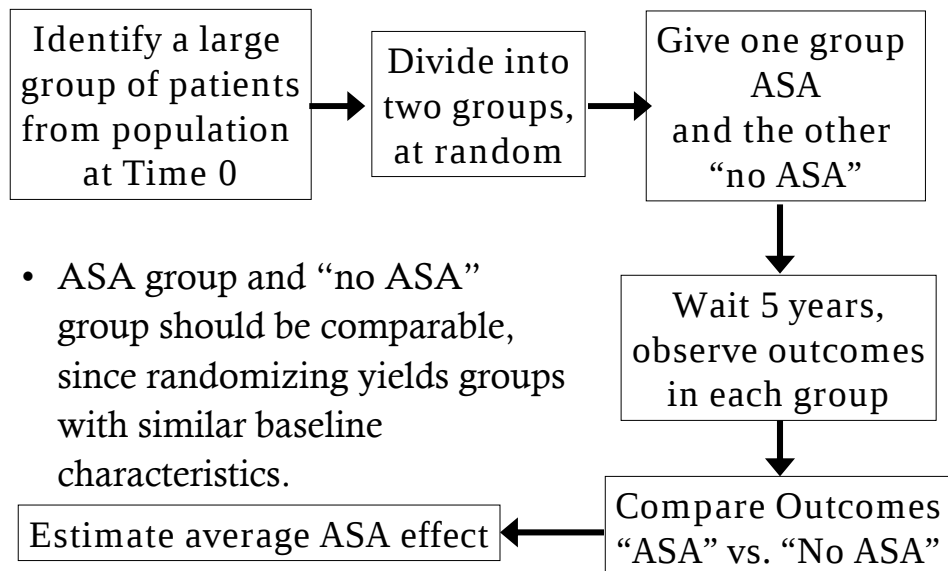
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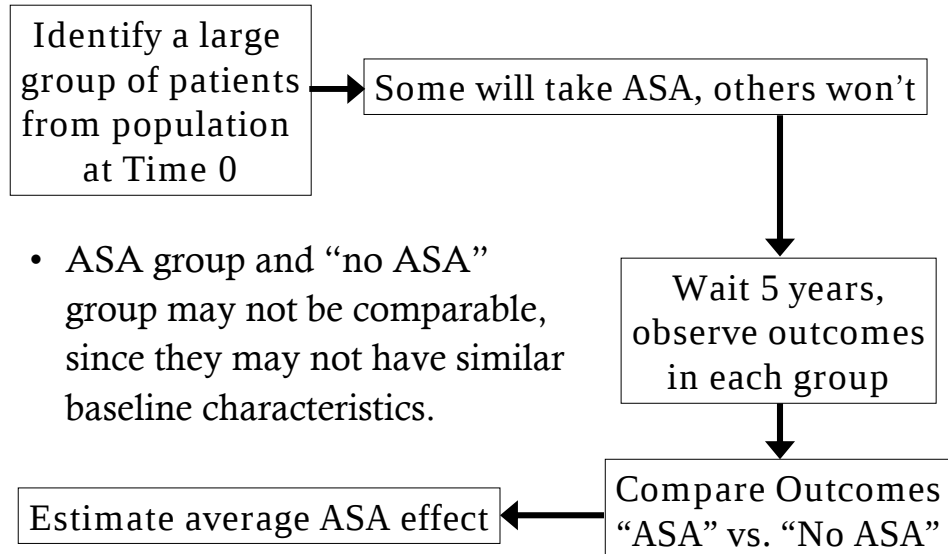
ASA & Mortality: Second Best Study



ASA & Mortality: RCT



ASA & Mortality: An Observational Study



Dealing With Selection Bias

- Analytical adjustments for Observed Bias
 - Matching, Stratification, Weighting, and Regression approaches are most common
- Goal: compare treated subjects to controls with same (or similar) baseline characteristics.
- Assuming no hidden bias (heroically), we use standard statistical comparisons of the groups after adjustment.

How Can We Avoid Being Misled?

1. What differentiates an observational study from a randomized controlled trial?
 - One key element: potential for selection bias.
2. What is selection bias, and why should I care about it?
 - Baseline characteristics of comparison groups are different in ways that affect the outcome.
3. What can be done to deal with selection bias in observational studies?
 - Propensity score methods for overt bias.
 - Sensitivity analyses to deal with hidden bias.

A Few Advantages of Propensity Methodology

- Results can be persuasive even to audiences with limited statistical training.
- Though estimating the PS requires some care, the comparability of treated and control patients can be verified simply.
- PS methods address selection bias well.
- PS methods may be combined with other sorts of adjustments.
- Methods appealing to grant review boards?

A Few Cautions and Limitations

- Hidden Bias: Beware unmeasured covariates which affect outcomes and/or treatment assignment.
- This is a reasonable method with fairly large samples.
- Options narrow as an investigation proceeds. What is easy early may become difficult or impossible later.
- Sadly, though OS work cries out for design, we're often working with secondary data, where a lot of design issues weren't considered in a serious way...

Strategic Issues in Observational Studies (Rosenbaum, 2002)

- Design observational studies
 - Exert as much experimental control as possible, carefully consider the selection process, and anticipate hidden biases
- Focus on simple comparisons
 - Increase impact of results on consumers
- Compare subjects who looked comparable prior to treatment
- Use sensitivity analyses to delimit discussions of hidden biases due to unobserved covariates

Rosenbaum (2002)

What should always be done in an OS ... and often isn't?

1. Collect data so as to be able to model selection
2. Demonstrate need for adjustment - selection bias
3. Carefully record intervention time - adjust only for things present before or at time of intervention.
4. Ensure baseline characteristic overlap [comparability]
5. Check baseline characteristic balance after adjustment
6. Specify relevant post-adjustment population with care
7. Estimate treatment effect in light of adjustment
8. Estimate sensitivity of results to potential hidden bias

Thank you very much.

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<http://www.chrp.org/propensity/ICHPR>

<http://www.chrp.org/love>

Rosenbaum (2002 text, Springer)

Rubin (2001, *Health Services & Outcomes Research Methodology*)

Austin & Mamdani (in press, *Statistics in Medicine*)

D'Agostino (1998, *Statistics in Medicine*)

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A Partially Annotated Bibliography

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A Partially Annotated Bibliography

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collected in three waves: confounders measured pre-treatment, while outcomes were measured post-treatment.]

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